Price indices with variable mark-ups and changing variety

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#### Idea

- · Price-index variations aim at measuring changes in the standard of living
- i.e. Money needed to obtain the same level of utility given new prices.
- Constructing a (micro-founded) price-index requires both prices, consumed quantities.
- · Prices and consumed quantities are influenced by,
  - Change in costs
  - Change in mark-ups
  - Change in demand (preferences)
  - Changes in variety
- Methodological contribution: propose a simple framework that encompasses all these effects.

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### The problem with estimating price-index changes

- Implicit to the previous definition is that preferences don't change
- and thus that the demand system does not change.
- Assumption easily violated by the data  $\longrightarrow$  so demand shocks are added (a.k.a. OLS residuals) ...
- ... which assume that preferences are changing.
- A way to solve this tension is proposed by Redding and Weinstein, 2018.

### Redding and Weinstein (2018): reinterpreting demand shocks

• Introduce preference (demand) shifters by introducing vector  $\varphi \in \mathbb{R}^n_{++}$ .

$$u(q) 
ightarrow u(\varphi q)$$
 where  $\varphi q = egin{bmatrix} arphi_1 \ arphi_2 q_2 \ arphi \ arphi_n q_n \end{bmatrix}$ 

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• Expenditure function,

$$e(p, u, \varphi) = \min_{q} p' q \text{ s.t. } u(\varphi q) \ge u,$$
  
=  $\min_{\tilde{q}} \left(\frac{p}{\varphi}\right)' \tilde{q} \text{ s.t. } u(\tilde{q}) \ge u, \qquad \text{set } (\tilde{q} = \varphi q)$   
=  $e(p/\varphi, u).$ 

#### Demand: introducing variable markups

Redding and Weinstein (2018) focus on CES-preferences

$$\mathcal{U} = \sum_{i=1}^{i=N} (\varphi_i q_i)^{rac{\sigma-1}{\sigma}}$$

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Leads to an iso-elastic demand when N is large
 → ill-equipped to capture variations in markups.

### Demand

• We assume instead a translog expenditure function as in Feenstra (2010) and Feenstra and Weinstein (2017).

$$\ln(e(p/\varphi)) = \alpha_0 + \alpha' \ln(p/\varphi) + \frac{1}{2} \ln(p/\varphi)' \Gamma \ln(p/\varphi).$$

with Σ<sup>n</sup><sub>i=1</sub> α<sub>i</sub> = 1, Γ is symmetric and each row/column sums to zero.
Then,

$$s_i(p/\varphi) = \alpha + \Gamma \ln(p/\varphi).$$

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### Price index

Exact price index given by an augmented Törnqvist index (Diewert, 1974)

$$\ln(e(p^{1}/\varphi^{1})) - \ln(e(p^{0}/\varphi^{0})) = \sum_{i=1}^{n} \frac{s_{i}^{1} + s_{i}^{2}}{2} (\ln(p_{i}^{1}/\varphi_{i}^{1}) - \ln(p_{i}^{0}/\varphi_{i}^{0})).$$

- Complication w.r.t. CES case: *n* is the total number of *potential* goods, some of which might not be available in a certain period.
- If *i* is not available then  $s_i = 0$  and  $p_i/\varphi^i$  is equal to the reservation price.
- We can't construct the Törnqvist as we do not observe shadow prices nor the demand shifters  $\varphi^i$ .

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### Shadow prices

- Reservation prices are a function of the prices of all available varieties through the demand system.
- Notations:
  - I set of common goods available in both periods;
  - $s_i^t$  expenditure share on *i* at time *t*;
  - $s_l^t$  total expenditure share on l at time t.
- New expression of our price-index:

$$\begin{split} \ln(e(p^{1}/\varphi^{1})) - \ln(e(p^{0}/\varphi^{0})) &= \frac{1}{\gamma} \sum_{i=1}^{n} \frac{(s_{i}^{0})^{2} - (s_{i}^{1})^{2}}{2} - \frac{1}{\gamma} \sum_{i \in I} \frac{(s_{i}^{0})^{2} - (s_{i}^{1})^{2}}{S_{i}^{0} + S_{i}^{1}}, \\ &+ \sum_{i \in I} \frac{s_{i}^{0} + s_{i}^{1}}{S_{i}^{0} + S_{i}^{1}} \left( \ln(p_{i}^{1}/\varphi_{i}^{1}) - \ln(p_{i}^{0}/\varphi_{i}^{0}) \right). \end{split}$$

where  $S_I^1 = \sum_{i \in I} s_i^1$  and  $S_I^0 = \sum_{i \in I} s_i^0$ .

- When there is no change in variety, this coincides with the augmented Törnqvist index.
- Still, we don't know  $\gamma$  and  $\varphi_i^1/\varphi_i^0$ .

#### Estimation of $\gamma$ - Price equation

- Oligopolistic competition.
- $\pi_i(p) = D_i(p)(p_i c_i)$  where  $D_i(p)$  is the demand for good *i*.
- Given the translog specifications, the FOC (taking expenditures as given) are,

$$p_i = c_i \left( 1 + \frac{s_i n}{\gamma(n-1)} \right),$$
$$\frac{p_i - c_i}{c_i} = \frac{s_i}{\gamma} \frac{n}{n-1} \approx \frac{s_i}{\gamma}.$$

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- Average mark-up is equal to  $\frac{1}{\gamma(n-1)}$ .
- Mark-up of firm *i* increases with its share and decreases with  $\gamma$ .

#### Estimation of $\gamma$ - Identifying assumption

Idea, independence between demand and supply shocks (Feenstra 1994)

$$(\varphi_i^1/\varphi_i^0) \perp (c_j^1/c_j^0)$$
 for all  $i, j \in I$ .

- Problem:  $\varphi_i^1/\varphi_i^0$  is not identified from data (given  $\gamma$ ).
- We make a "difference in difference" (time periods + goods)
- · With our pricing equation, the identifying assumption is

$$\theta = E\left[\left(\ln(\varphi_i^1/\varphi_i^0) - \ln(\varphi_j^1/\varphi_j^0)\right)\left(\ln(c_i^1/c_i^0) - \ln(c_j^1/c_j^0)\right)\right] = 0.$$

• Holds if  $cov(\ln(\varphi_i^1/\varphi_i^0), \ln(c_j^1/c_j^0)) = 0$  for  $i, j \in I$ .

### Estimation

• Finite sample moment condition,

$$\hat{\theta}_{|I|} = \frac{1}{|I|(|I|-1)} \sum_{i,j \in I} \left[ \left( \ln(\varphi_i^1/\varphi_i^0) - \ln(\varphi_j^1/\varphi_j^0) \right) \left( \ln(c_i^1/c_i^0) - \ln(c_j^1/c_j^0) \right) \right] = 0$$

• This is a U-statistic (van der Vaart, 1998),

$$\sqrt{|I|} (\hat{ heta}_{|I|} - heta) \sim^{a} \mathcal{N}(0, 4\sigma^{2}),$$

where  $\sigma^2$  is the variance of,

$$\mathsf{E}\left[\left(\mathsf{ln}(\varphi_i^1/\varphi_i^0) - \mathsf{ln}(\varphi_j^1/\varphi_j^0)\right) \left(\mathsf{ln}(c_i^1/c_i^0) - \mathsf{ln}(c_j^1/c_j^0)\right) \middle| \varphi_i^1/\varphi_i^0; c_i^1/c_i^0\right].$$

- We estimate  $(1/\gamma)$  by minimizing  $(\hat{\theta}_{|I|})^2$  (minimum should be zero).
- $\sigma^2$  can be consistently estimated using the finite sample analogue and the estimate of  $1/\gamma$ .
- Confidence intervals for  $\widehat{1/\gamma}$  are derived using the Delta method.

## Validity check

Monte Carlo simulation (1000 draws).

n	1	rejection rate	
		(nominal level 5%)	
30	10	0.164	
90	30	0.087	
150	50	0.068	
300	100	0.060	
900	300	0.056.	

## Application/Illustration

- ACNielsen homescan panel for Denver area (Aguiar and Hurst, 2007).
- Grocery packed goods purchases for a large number of households from Jan 1993 until March 1995
- Focus on 1993/1994 price indices
- We aggregate over consumers and year. Total value purchased of every good and total quantity purchased of every good.
- Good = UPC code
- Analysis for each product sub-category: 48 in total.
  - example: apple juice (42 upc's); orange juice (60 upc); tomato sauce (131 upc); soup (308 upc); instant (noodles) (157 UPC)...
  - We focus on groups with more than 30 goods common for both periods: 30 groups

## Application/Illustration

#### • Summary statistics over the 30 groups.

	mean	std	min	max
$\gamma$	0.083257	0.070692	0.012714	0.31657
mean mark-up	0.25797	0.098305	0.098714	0.54176
/	93.7	78.480	30	321

### Price decomposition

• Price decomposition

$$\begin{split} &\ln(e(p^1/\varphi^1)) - \ln(e(p^0/\varphi^0)), \\ &= \frac{1}{\gamma} \sum_{i=1}^n \frac{(s_i^0)^2 - (s_i^1)^2}{2} - \frac{1}{\gamma} \sum_{i \in I} \frac{(s_i^0)^2 - (s_i^1)^2}{S_I^0 + S_I^1}, & \text{variety effect} \\ &+ \sum_{i \in I} \frac{s_i^0 + s_i^1}{S_I^0 + S_I^1} \ln((p_i^1/c_i^1)/(p_i^0/c_i^0)), & \text{mark-up effect} \\ &+ \sum_{i \in I} \frac{s_i^0 + s_i^1}{S_I^0 + S_I^1} \ln(c_i^1/c_i^0), & \text{cost-shift effect} \\ &- \sum_{i \in I} \frac{s_i^0 + s_i^1}{S_I^0 + S_I^1} \ln(\varphi_i^1/\varphi_i^0). & \text{demand-shift effect} \end{split}$$

Given (the estimate of)  $\gamma$  we can identify the first three effects. The fourth part (demand effect) is not identified without additional assumptions.

• We impose that,

$$\sum_{i \in I} \ln(\varphi_i^1 / \varphi_i^0) = 0, \text{ then },$$

$$\ln(\varphi_i^1 / \varphi_i^0) = \frac{s_i^1 - s_i^0}{\gamma} - \frac{S_I^1 - S_i^0}{\gamma |I|} + \ln(p_i^1 / p_i^0) - \sum_{i \in I} \frac{\ln(p_i^1 / p_i^0)}{|I|}.$$

# Example



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## Price decomposition

#### • Summary statistics (in %-points)

	varieties	mark-up	supply	demand	total
mean	-0.64418	-0.74634	0.33683	0.36598	-0.86497
std	1.0592	1.7483	6.3052	3.9010	4.5696
min	-3.5287	-4.8344	-9.9060	-15.8201	-11.3404
max	1.3733	1.8282	24.7810	5.1179	7.8084

• Correlation matrix (in %-points)

	varieties	mark-up	supply	demand	total
varieties	1.000	0.932	0.056	-0.460	0.235
mark-up	0.932	1.000	-0.005	-0.507	0.125
supply	0.056	-0.005	1.000	-0.581	0.782
demand	-0.460	-0.507	-0.581	1.000	-0.140
total	0.235	0.125	0.782	-0.140	1.000

# Conclusion

- Importance of mark-ups
- Decomposition of price index into different components
- U-statistic based estimator for  $\gamma$  (for variable number of goods).
- Future research
  - Belgian supermarket data (?)
  - multi-product firms
  - Trade data (mulitple years/countries of origin)
  - Comparison with CES, including a supply-free for identification of  $\gamma$

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Beyond translog