

# Climate Policies and the Carbon Content of Jobs

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# Outline of the presentation

Background

Carbon content of jobs

Facts about carbon-intensive occupations

Empirical strategy

Results

Discussion

# Climate policies create winners and losers

- **Distributional effects in the labour markets have large echo in the political debate:**
  - ◇ President Biden: “When I hear **climate**, I think jobs, **good-paying union jobs...**”
  - ◇ Congresswoman Bachmann renamed the **Environmental Protection Agency** “**the job-killing organization of America.**”
- Such polarized debate **obscures** the key issues to design fair green policy packages:
  - ◇ In macro models aggregated employment effects usually small → Several papers understate the problems associated with the transition (e.g., Metcalf, 2023).
  - ◇ But **distributional effects large** for certain groups → Ensuring a smooth reallocation for displaced workers requires identifying carbon-intensive jobs beyond coal miners.
  - ◇ As for trade shocks, **unmanaged distributional effects** in the labour market fuel **political opposition** against climate policies (e.g., Vona, 2023).

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# This paper in a nutshell

- Using very granular data on **CO2 emissions**, we build a **time-varying measure of carbon intensity** for **411 occupations** over the period **2003-2019**.
- The **carbon content** of an occupation (i.e., a weighted average of the establishment/sector CO2 intensity) is used to **evaluate** the impact of two climate policies: **energy prices (today)** and **the EU-ETS (just preliminary results)**.
- We empirically show that such measure:
  - ◊ is a better proxy of **vulnerability** than a sector measure and allows to compute the **share of workers at risk of displacement**;
  - ◊ sheds light on **heterogeneous effects** of a climate policy, especially on labour demand for different occupations;
  - ◊ highlights that **low- and middle skilled carbon-intensive occupations** are **particularly affected** by energy price shocks.

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# Occupation vs. sector: evaluation of climate policies

Existing literature on environmental economics uses **sector-level measures** of exposure to estimate labour market impacts.

- The **Clean Air Act** quasi-experiment, triple-difference approach: job losses concentrated in polluting industries (Greenstone, 2002; Walker, 2011; Curtis, 2017).
  - **Energy price effects**, either using shift-share instruments (Marin and Vona, 2019, 2021) or border-pair fixed effects (Kahn and Mansur, 2013): job losses significantly larger in energy-intensive sectors.
  - **EU-ETS papers** combine matching and DID: no clear negative effects on employment (Marin et al., 2018; Dechezlepretre et al., 2020; Colmer et al., 2022).
- ⇒ **Issue:** using sectoral exposure, all occupations within a carbon-intensive sector are assumed to be equally exposed to carbon pricing shocks.

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# Occupational heterogeneity & climate policies

A few recent papers suggest that **policy effects** are **heterogeneous** across **skill/occupational groups** within the same sector.

- The employment impact of the **British Columbia carbon tax** negligible on aggregate (Yamazaki, 2017), but **strongly negative** on **low-skilled workers** (Yip, 2019).
- Marin and Vona (2019, 2021) show that **employment losses** induced by **energy prices** are **highly heterogeneous** across **occupations**, also within the same sector.
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# Occupation vs. sector to measure exposure to shocks

Research in labour economics shows that occupations convey more information than sectors along several dimensions:

- **Skills (both general and specific)**
  - ◇ Kambourov and Manovskii (2008): human capital specificity resides in occupational rather than industry categories,
  - ◇ Poletaev and Robinson (2008): occupation-specific skill proximity key predictor of post-displacement earnings.
- **Exposure to structural shocks**
  - ◇ Skill-biased technical change literature: within-sector across occupation effects dominant,
  - ◇ Autor et al. (2003), Goos et al. (2014): routine task intensity indicator key predictor of labour market outcomes.
- **Bargaining power and outside option of workers**
  - ◇ Acemoglu et al. (2001): skill-biased shocks → bargaining more decentralized at the skill-level,
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# Measuring occupational exposure to climate policies

However, several questions remain unanswered in these recent papers:

- Is it possible to build a **single, continuous and time-varying** occupation-based measure of **vulnerability** to climate policies similar of those of Frey and Osborne (2013) or Autor et al. (2003) for digital techs?
- To what extent **existing results** on energy prices and EU-ETS impacts **hide** substantial **heterogeneity** across occupations? Which **dimension** of **worker's heterogeneity** is the most important?
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# Contributions

We contribute to three strands of literature:

1. **Just transition**  $\Rightarrow$  providing a new **measure of vulnerability** that can be implemented with **less granular** data and highlighting **new profiles of vulnerable workers** beyond coal miners;
2. Labour market impacts of structural transformations:
  - $\diamond$  building a time-varying measure that incorporates carbon-saving technological change;
  - $\diamond$  comparing exposure to climate policies with exposure to other structural shocks;
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## Description of data

Our main analysis is at **occupation-by-establishment** level.

- **Unbalanced panel of establishments for 1997-2019** from **EACEI** (Enquête Annuelle sur le Consommations d'Énergie dans l'Industrie), with **unit** of analysis the **establishment** (SIRET).
  - ◊ **Survey on consumption and expenditure** for energy products (by **source**: electricity, oil, coal, gas, steam, other)
  - ◊ **Stratified sample** of medium-small **manufacturing** establishments (10-250 employees) and population of big manufacturing establishments (250+ employees)
- **Balanced panel of 411 occupations** from **DADS** (Déclaration Annuelle des Données Sociales), with **unit** of analysis the **establishment** (SIRET), only from 2003
  - ◊ DADS contains occupational employment shares and wages for the **universe** of French establishments → very accurate measures
  - ◊ Information on **employment** (in FTE) and **annual wages** by **occupation** (PCS)
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## New metrics: carbon-content of occupations

- **Binary definition of brown jobs** fails to capture carbon-reducing technological change and **different degrees** of vulnerability beyond coal miners.
- We build an index of the carbon content of occupations for 400+ occupations over the period 2003-2018 capturing the worker's outside option to carbon pricing shocks:

$$CC_{ot} = \sum_{i=1}^N \frac{L_{oit}}{L_{ot}} \times \frac{CO2_{it}}{L_{it}},$$

where  $i$  indexes EACEI establishments in manufacturing and 3-digit industries for non-manufacturing sectors (JRC-Eurostat data) or manufacturing establishments not surveyed in EACEI

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## Theoretical interpretation of the carbon-content

- **Carbon-task specificity**, the demand of occupations and tasks specific to **high-carbon productions** will decline → we expect **larger employment losses** for high-carbon occupations, less clear is **by how much**.
- **Outside option and wage effects**, for a worker employed in sector  $j$ , the carbon content can be approximated as:

$$CC_{ot} \approx \underbrace{\frac{CO2_{jt}}{L_{jt}}}_{\text{actual exposure}} + \underbrace{\sum_{i \neq j}^{N-1} \frac{L_{oit}}{L_{ot}} \frac{CO2_{it}}{L_{it}}}_{\text{outside option}},$$

- ◊ A higher carbon content reveals a **weaker bargaining position** to carbon pricing shocks.
- ◊ But the effect of **carbon pricing on firm profits and workers' quasi-rents** is unclear (especially for the EU-ETS).
- ◊ **Workers' selection** may increase the average workers' skills and thus wages in firms more exposed to carbon price shocks.

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# Greenness indicator

- For descriptive purposes, we compare labour market dynamics of **carbon intensive** and **green occupations**.
- Carbon intensity and the greenness of an occupation capture **two different aspects** of the labour market adjustment: **vulnerability vs. reducing env. impact**.
- Building on our previous work (Vona *et al.*, 2015, 2018, 2019), we use a **task-based indicator** of 'greenness':

$$Greenness_k = \frac{\# \text{ green tasks}_k}{\# \text{ tasks}_k}$$

- The greenness captures the **relative importance** (e.g. time spent) of **green tasks** for that occupation.
- Data on green tasks are only available for the US, thus we use a **cross-walk** of US SOC 6-digit occupations to French PCS 4-digit occupations to retrieve the occupational greenness.

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## Climate policy measures: energy prices

- As in Davis et al (2013), what we call **energy price** is, actually, the average **unit cost of energy**, i.e. expenditures divided by quantity consumed (in kWh)
- This **ratio** can be written as:

$$P_{et}^E = \sum_{j=1}^J \phi_{et}^j P_{et}^j,$$

where  $\phi_{et}^j$  is the **share** of **energy** consumption of source  $j$  (i.e. gas, electr, coal, oil, etc) on **total** energy consumption, while  $P_{et}^j$  is the **price** of energy source  $j$  paid by **establishment**  $e$  at time  $t$

- Similar to Jo (2022), here we distinguish between the **price of dirty** (i.e., all fossil fuels) and **clean energy** (e.g., electricity in France) → only the former should have an impact on the carbon content.

# Facts about carbon-intensive occupations

1. We observe no **“unconditional” catching-up** (hard to decarbonize?), but a mild **“conditional” catching-up** in the carbon content of occupations;
2. Carbon intensive occupations are **more vulnerable in general**, being more exposed to other **skill-biased** structural shocks;
3. Carbon-intensive occupations exhibit **slower employment growth**;
4. **Wage growth** is uncorrelated (or slightly positively correlated) with carbon intensity;
5. Little overlapping between **greenness** and **carbon-intensity** of jobs. The two groups exhibit **opposite patterns** in terms of employment and wages.

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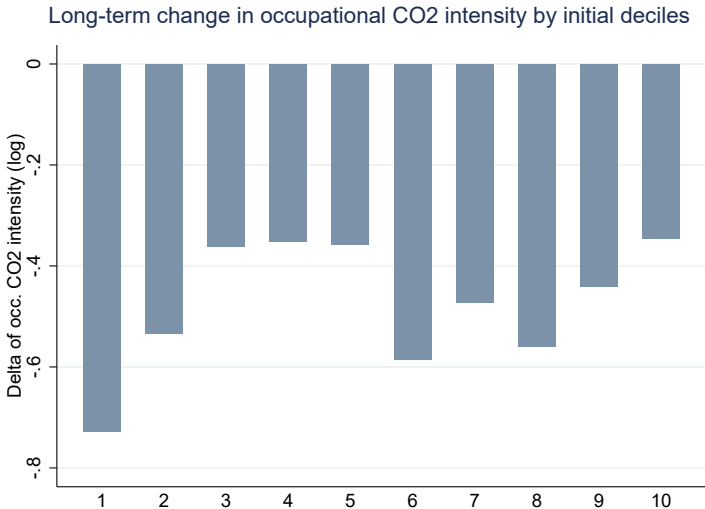
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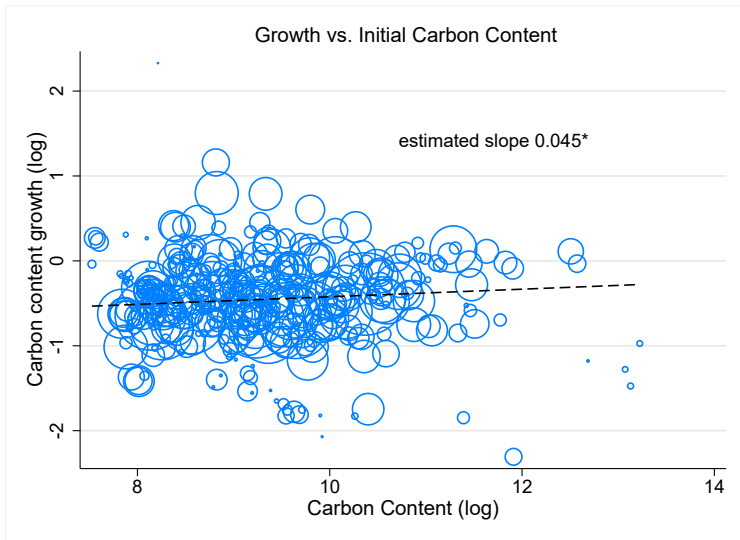


# FACT I: Unconditional catching-up



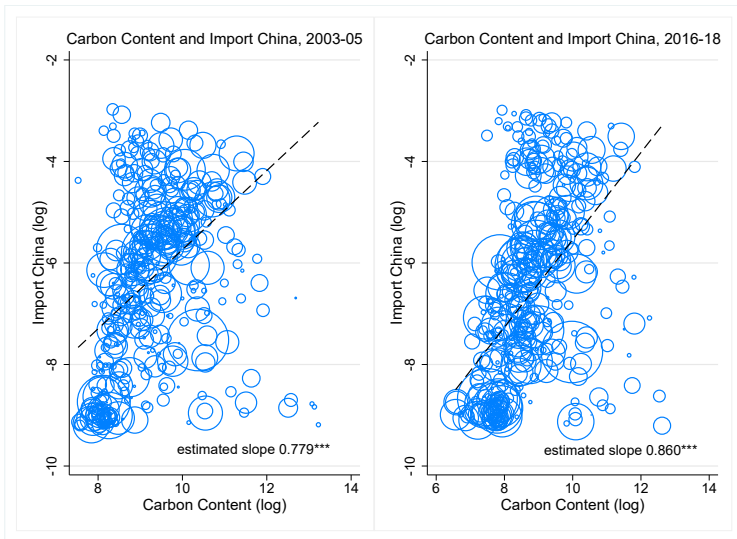
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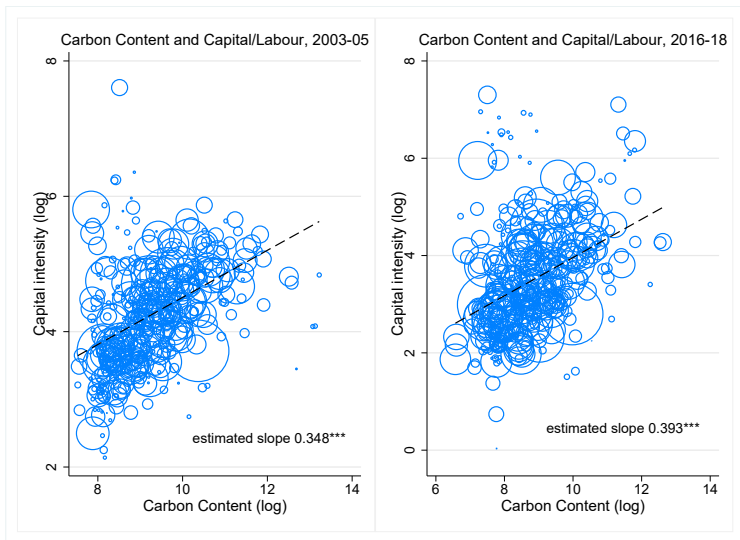
source: our elaborations on EACEI-DADS-JRC data.

## FACT II: Carbon content and trade



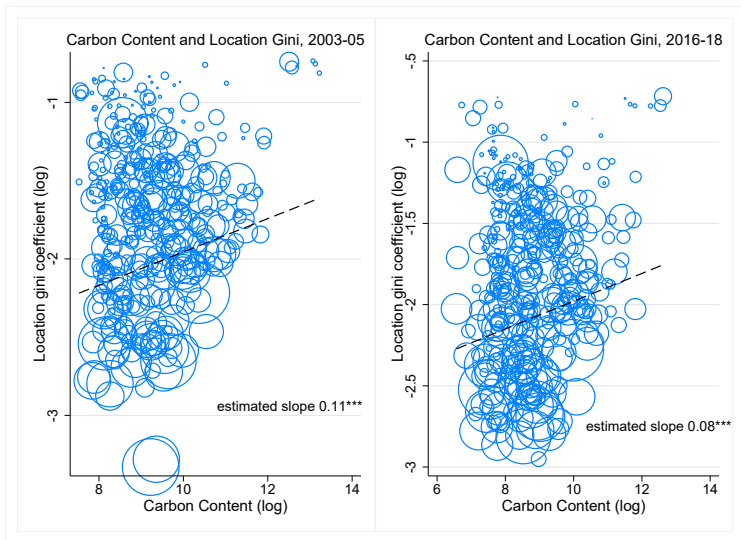
source: our elaborations on EACEI-DADS-JRC data. Slopes:  $\beta_{init} = 0.289$ ,  $\beta_{end} = 0.418$

## FACT II: Carbon content and capital deepening



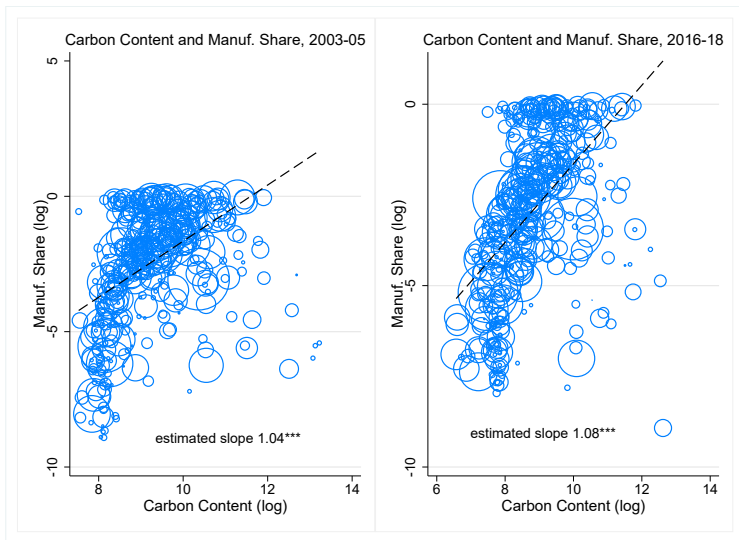
source: our elaborations on EACEI-DADS-JRC data. Slopes:  $\beta_{init} = 0.313$ ,  $\beta_{end} = 0.363$

## FACT II: Carbon content and spatial concentration



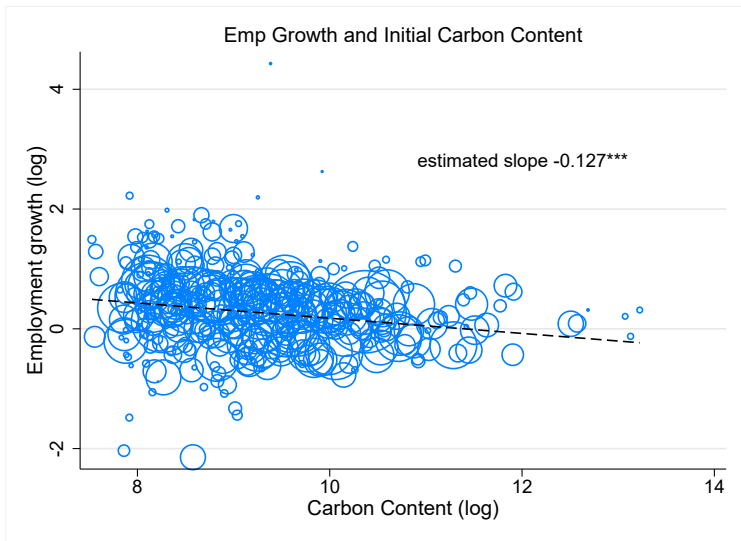
source: our elaborations on EACEI-DADS-JRC data. Slopes:  $\beta_{init} = 0.107$ ,  $\beta_{end} = 0.113$

## FACT II: Carbon content and Manuf. share



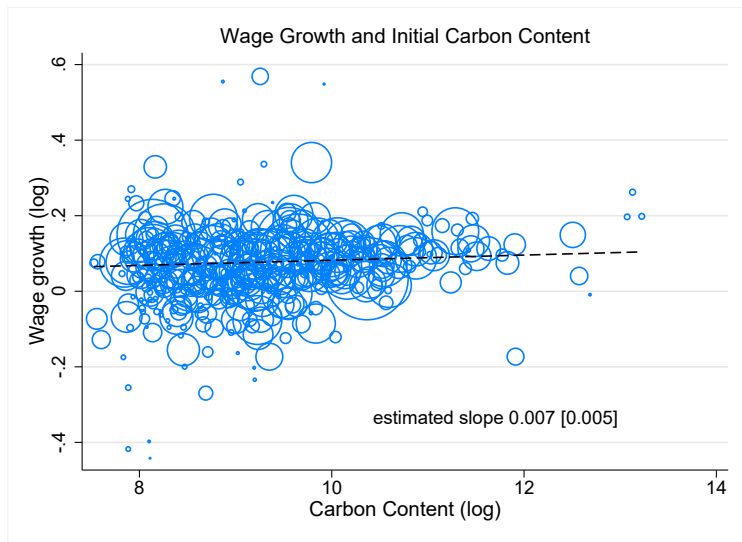
source: our elaborations on EACEI-DADS-JRC data. Slopes:  $\beta_{init} = 1.144$ ,  $\beta_{end} = 1.022$

## FACT III: Initial carbon content and emp. growth



Growth-growth spec.: similar results [here](#)

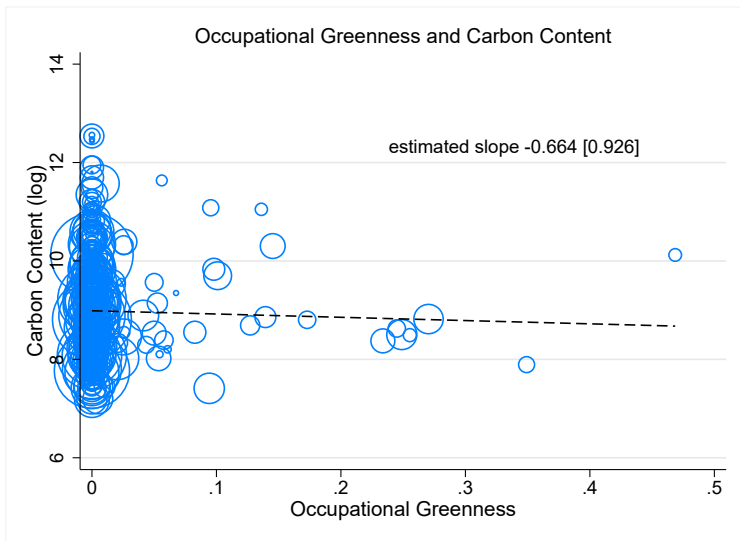
## FACT IV: Initial carbon content and wage growth



Growth-growth spec.: similar results [here](#)



# FACT V: Green vs. carbon-intensive occupations



source: our elaborations on EACEI-DADS-JRC data

## Conditional correlations

We estimate the relationship btw **wage/employment/CO2 intensity** and occupational **initial CO2 intensity** and **greenness**, controlling for: capital intensity, imports from China, etc. and occ. FE.

$$\log(y_{ot}) = \beta_1 \log(CC_{ot}) + \beta_2 \text{greenness}_o \times t + \gamma X'_{ot} + \mu_t (+\mu_{ot}) + \epsilon_{ot}$$

Table: Conditional correlations

Dep. var:	log(CO2/L)	log(FTE)	log(wages)
init. log(CO2/L) x time	-0.0292** (0.0119)		
log(CO2/L)		-0.083** (0.030)	0.0054 (0.0036)
Greenness x time	-0.2851 (0.2598)	0.203** (0.097)	-0.104*** (0.025)
Controls: other shocks	Yes	Yes	Yes
2-digit occ. x years F.E.	Yes	Yes	Yes
R sq	0.340	0.501	0.901
N of 4-digit occupations	411	411	411
N	2,055	2,055	2,055

Notes: FE estimates weighted by initial occ. employment FTE. Control variables: see above. Standard errors clustered by 4-digit PCS occupation in parenthesis. \* p<0.1, \*\* p<0.05, \*\*\* p<0.01.

## Estimating energy price impacts

We estimate the following equation for an unbalanced panel of establishment  $e$ -occupation  $o$  pairs for the period 2003-2018:

$$\log(Y_{oet}) = \beta_1 \log(P_{et}^E) + \beta_2 \log(CC_{ot}) + \beta_3 \log(P_{et}^E) \times \log(CC_{ot}) + \dots \\ \alpha_{eo} + \xi_{st} + \phi_{rt} + \gamma_t \mathbb{1}_{k \in ETS(t)} + \eta_{ot} + X'_{ot} \mu + \varphi \text{green}_o \times t + \varepsilon_{oet}$$

where:

- $Y_{oet}$  is FTE employment/wages (in log);
- $P_{et}^E$  is the average price of energy in establishment  $e$ ;
- $CC_{ot}$  is the CO2 emission intensity in occupation  $o$ .
- **Favourite** specification controls for **fixed effects**: estab.-occ. ( $\alpha_{eo}$ ), sector(2-digit NACE)-by-year ( $\xi_{st}$ ), region-by-year ( $\phi_{rt}$ ) and occupation-by-year ( $\eta_{ot}$ ) as well as EU-ETS-by-year dummies ( $\mathbb{1}_{k \in ETS(t)}$ ).
- **Main source of identifying variation**: within-estab.-occ. effects of energy price shocks net of sector-, region-, occupation- and ETS-specific trends.

# Challenges for the empirical analysis

- **Zero inflation:** several establishment-by-occupation cells are zeros.
  - ◇ **Self-selection** problems affecting average occupational wages → we keep only **occupations** present in the **initial period** for **wages**.
  - ◇ For **employment:** flexible log-transformation to account for the zeros ( $\log(x + \text{min}/2)$ ) and analysis at both the 4digit occupation vs. 2digit occupation level, where less zeros.
- **Potential endogeneity of the occupational carbon content** → we use the sector (3digit) CO2 intensity to build a measure of the carbon content of the occupation.
- **Testing sector- vs. occupation- effects:** which one is prevalent? Recall that  $CC_{ot}$  decomposed in two components: sector of work and outside option.

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## Challenges for the empirical analysis (cnt.)

- **Endogeneity** of energy prices → two main sources (Marin and Vona, 2021):
  - ◇ quantity discounts: larger firms pay lower prices;
  - ◇ unobserved technological change: L-E substitution vs. accelerating automation (K replaces both E and L).
- **Dirty and clean fuels** → fossil fuels vs. electricity → we use **only dirty fuel prices** in the main specification controlling for **initial electricity share** of the establishment interacted with year dummies.
- For the **EU-ETS quasi-experiment**: as common practice in the literature, we combine **difference-in-difference** and **matching** to retrieve the causal effect of the EU-ETS as a function of the occupational carbon content.

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## Instrumental variable

- As in Marin and Vona (2021), we build a shift-share **IV** that **only** keeps **exogenous variations** in energy **prices** and **accounts** for **both sources of endogeneity**

$$P_{et}^{IV} = \sum_{j=1}^J \phi_{e,t=\text{presample}}^j P_t^j$$

- Domestic regulation induced substantial changes in prices, especially for electricity;
- Prices for other sources respond more to 'global' prices;
- We shut down endogenous responses of establishments to changing energy prices by weighting exogenous prices with a time-invariant (lagged) establishment-specific energy mix
- Energy mix observed in the first year available in EACEI, lagged at least 3 years from the first observation.

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## Endogeneity of energy prices (cnt.)

- We adapt **recent insights** on shift-share IV design to enhance the credibility of identification strategy,
- Note that the approach of treating energy price shocks **as-good-as randomly assigned** conditional on **X** applies if the number of shocks is **large** (Borusyak et al., 2021).
- Here only **four fuels** account for the bulk of energy consumed (electricity, gas, heating oil, coal) [here](#)
- Thus, a Bartik instrument is equivalent to use initial **local shares** (i.e. **energy source shares**) as instruments (Goldsmith-Pinkham et al., 2020):
  - ◊ We test the **parallel trends** assumption with respect to energy source shares. [here](#)
  - ◊ We account for **unobserved heterogeneity** in levels using **estab.-occ. fixed effects**.
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## Employment, 2digit, all fuels

To ease the interpretation, the occupational carbon-intensity is net of the sample median, some stats: IQR=0.703;  $p(90)-p(50)=1.39$ . Median of the  $\log(\text{energy} - \text{price}) = 4.05$ ,  $\log(\text{dirty} - \text{price}) = 3.73$ . The number of 2digit occ. is 29.

**Table:** Results for employment (2-digit occ.) - average energy price

Dep var: FTE empl (log)	(1) FE	(2) FE-IV	(3) FE	(4) FE-IV	(5) FE	(6) FE-IV
Energy price (log)	-0.241*** (0.0099)	-0.112* (0.0584)	-0.058*** (0.0068)	0.0023 (0.0573)	-0.042*** (0.0069)	0.0035 (0.0559)
Carbon content of occ. (log)	1.583*** (0.0070)	1.598*** (0.0072)	-0.045*** (0.0075)	-0.027*** (0.0081)		
Energy price x Carbon cont. of occupations	-0.492*** (0.0145)	-0.446*** (0.0235)	-0.084*** (0.0063)	-0.180*** (0.0117)	-0.045*** (0.0068)	-0.189*** (0.0369)
Reg.-year FE	✓	✓	✓	✓	✓	✓
Sect.-year FE	✓	✓	✓	✓	✓	✓
Estab. FE	✓	✓				
Occ.-estab. FE			✓	✓	✓	✓
Occ.(2-digit)-year FE					✓	✓
F test of excluded IV		277.2		276.8		259.5
N of establishments	13327	13327	13120	13120	13120	13120
N. obs	1206166	1206166	1204053	1204053	1204053	1204053

Notes: Unit of analysis: establishment-occupation (2-digit PCS) pair. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Additional control: ETS-year dummies.

# Employment, 2digit, dirty price

**Table:** Results for employment (2-digit occ.) - average price of fossil fuels

Dep var: FTE empl (log)	(1) FE	(2) FE-IV	(3) FE	(4) FE-IV	(5) FE	(6) FE-IV
Dirty-energy price (log)	-0.173*** (0.0082)	-0.315*** (0.0648)	-0.038*** (0.0051)	-0.069 (0.0620)	-0.022*** (0.0051)	-0.090 (0.0581)
Carbon content of occ. (log)	1.632*** (0.0073)	1.665*** (0.0076)	-0.040*** (0.0079)	-0.021** (0.0087)		
Dirty price x Carbon cont. of occupations	-0.335*** (0.0140)	-0.482*** (0.0266)	-0.052*** (0.00489)	-0.137*** (0.0102)	-0.019*** (0.00506)	-0.181*** (0.0458)
Decile electr. share-year FE	✓	✓	✓	✓	✓	✓
Reg.-year FE	✓	✓	✓	✓	✓	✓
Sect.-year FE	✓	✓	✓	✓	✓	✓
Estab. FE	✓	✓				
Occ.-estab. FE			✓	✓	✓	✓
Occ.(2-digit)-year FE					✓	✓
F test of excluded IV		134.0		134.1		138.6
N of establishments	10398	10398	10398	10398	10398	10398
N. obs	1004495	1004495	1004495	1004495	1004495	1004495

Notes: Unit of analysis: establishment-occupation (2-digit PCS) pair. Standard errors clustered by establishment in parenthesis. \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Additional control: ETS-year dummies.

# Employment, 2digit, dirty, multiple exposure

**Table:** Results for employment (2-digit occupations) - overlapping exposures

Dep var: FTE empl (log)	(1) FE	(2) FE-IV	(3) FE	(4) FE-IV
Dirty-energy price (log)	-0.0169*** (0.0061)	-0.0976 (0.0952)	-0.0159*** (0.0058)	-0.0595 (0.0714)
Dirty-energy price x Occ.	-0.0186*** (0.0051)	-0.182*** (0.0449)	-0.0147** (0.0064)	-0.152*** (0.0491)
Carbon content (log)				
Dirty-energy price x Sect.	-0.00451 (0.0029)	0.00481 (0.0400)		
Carbon content (log)				
Dirty-energy price x Occ.			-0.0083 (0.0671)	-0.774 (0.543)
Gini location coeff.				
Dirty-energy price x Occ.			0.0574* (0.0328)	0.407 (0.352)
import penetration				
Dirty-energy price x Occ.			-0.0285*** (0.0095)	-0.110** (0.0448)
capital intensity (log)				
F test of excluded IV		36.67		57.58
N of establishments	10398	10398	10398	10398
N. obs	1004495	1004495	1004495	1004495

Notes: Unit of analysis: establishment-occupation (2-digit PCS) pair. Standard errors clustered by establishment in parenthesis. \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Additional controls in all specification: occupation-establishment fixed effects, occupation-year dummies, sector-year dummies, region-year dummies, ETS-year dummies, initial decile of electricity share x year dummies.



# Employment, 4digit, main results

**Table:** Results for employment (4-digit occupations)

Dep var: FTE empl (log)	(1)	(2)	(3)
Energy price (log)	0.0613 (0.0386)		
Energy price (log) x Carbon cont. of occupation (log)	-0.0874*** (0.0145)		
Dirty-energy price (log)		-0.0242 (0.0385)	-0.0947 (0.0713)
Dirty-energy price (log) x Carbon cont. of occupations (log)		-0.0414** (0.0172)	-0.0460*** (0.0166)
Dirty-energy price (log) x Sectoral emission intensity (log)			0.0427 (0.0275)
F test of excluded IV	202.1	114.2	28.58
N of establishments	13048	10345	10345
N. obs	5429943	4701391	4701391

Notes: Unit of analysis: establishment-occupation (4-digit PCS) pair. IV-FE model. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Additional controls in all specification: occupation-establishment fixed effects, occupation-year dummies, sector-year dummies, region-year dummies, ETS-year dummies, initial decile of electricity share x year dummies (except column 1).

# Employment, 4digit, by occ.

**Table:** Results for employment (4-digit occupations) - by occupational skill level

Dep var: FTE empl (log)	Low-skill occ	Medium- skill occ	High-skill occ
Dirty-energy price (log)	-0.0536 (0.0644)	-0.0272 (0.0442)	-0.00106 (0.0599)
Dirty-energy price (log) x Carbon cont. of occupations (log)	-0.100*** (0.0354)	-0.0464** (0.0193)	0.0327 (0.0329)
F test of excluded IV	102.3	117.4	102.0
N of establishments	10028	10343	10190
N. obs	689591	2871974	1139826

Notes: Unit of analysis: establishment-occupation (4-digit PCS) pair. IV-FE model. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Additional controls in all specification: occupation-establishment fixed effects, occupation-year dummies, sector-year dummies, region-year dummies, ETS-year dummies, initial decile of electricity share x year dummies.

# Wage, 2digit, all fuels

**Table:** Results for wages (2-digit occupations) - average energy price (all energy inputs)

Dep var: hourly wage(log)	(1) FE	(2) FE-IV	(3) FE	(4) FE-IV	(5) FE	(6) FE-IV
Energy price (log)	-0.022*** (0.0031)	-0.0326 (0.0257)	0.0032 (0.0023)	-0.037** (0.0183)	-0.0025 (0.0023)	-0.052*** (0.0172)
Carbon content occ. (log)	-0.069*** (0.0023)	-0.071*** (0.0024)	0.072*** (0.0037)	0.067*** (0.0039)		
Energy price x Carb. cont. of occupation (log)	-0.073*** (0.005)	-0.054*** (0.0081)	0.014*** (0.0029)	0.042*** (0.0056)	-0.010*** (0.0036)	-0.0198 (0.0177)
Reg.-year FE	✓	✓	✓	✓	✓	✓
Sect.-year FE	✓	✓	✓	✓	✓	✓
Estab. FE	✓	✓				
Occ.-estab. FE			✓	✓	✓	✓
Occ.(2-digit)-year FE					✓	✓
F test of excluded IV		289.4		288.8		267.0
N of establishments	13116	13116	13116	13116	13116	13116
N. obs	849954	849954	849954	849954	849954	849954

Notes: Unit of analysis: establishment-occupation (2-digit PCS) pair. Standard errors clustered by establishment in parenthesis. \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Additional controls in all specification: sector-year dummies, region-year dummies, ETS-year dummies.

# Wage, 2digit, dirty

**Table:** Results for wages (2-digit occupations) - average price of fossil fuels

Dep var: hourly wage (log)	(1) FE	(2) FE-IV	(3) FE	(4) FE-IV	(5) FE	(6) FE-IV
Dirty-energy price (log)	-0.016*** (0.0024)	-0.059** (0.0277)	0.004*** (0.0016)	-0.021 (0.0185)	0.0004 (0.0016)	-0.033** (0.0153)
Carbon content occ. (log)	-0.064*** (0.0024)	-0.064*** (0.0026)	0.072*** (0.0039)	0.067*** (0.0042)		
Dirty price x Carb. cont. of occupation (log)	-0.054*** (0.0047)	-0.067*** (0.0094)	0.0095*** (0.0024)	0.033*** (0.005)	-0.0045 (0.0028)	0.0156 (0.0243)
Reg.-year FE	✓	✓	✓	✓	✓	✓
Sect.-year FE	✓	✓	✓	✓	✓	✓
Estab. FE	✓	✓				
Occ.-estab. FE			✓	✓	✓	✓
Occ.(2-digit)-year FE					✓	✓
F test of excluded IV		151.2		156.2		161.9
N of establishments	10854	10854	10854	10854	10854	10854
N. obs	732991	732991	732991	732991	732991	732991

Notes: Unit of analysis: establishment-occupation (2-digit PCS) pair. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Additional controls in all specification: sector-year dummies, region-year dummies, ETS-year dummies, initial decile of electricity share x year dummies.

# Wage, 2digit, dirty, multiple exp.

**Table:** Results for wages (2-digit occupations) - overlapping exposures

Dep var: hourly wage (log)	(1) FE	(2) FE-IV	(3) FE	(4) FE-IV
Dirty-energy price (log)	-0.0009 (0.0016)	-0.047* (0.0246)	0.000 (0.0016)	-0.043*** (0.0159)
Dirty price x Carbon content of occupation (log)	-0.005* (0.0028)	0.008 (0.0109)	-0.0015 (0.0039)	0.0266 (0.0278)
Dirty-energy price (log) x Sect. emission intensity (log)	0.0012 (0.0011)	0.0143 (0.0239)		
Dirty price x Occ. Gini location coefficient			-0.0078 (0.0303)	-0.230 (0.236)
Dirty price x Occ. import penetration			-0.0091 (0.0103)	-0.0644 (0.0889)
Dirty price x Occ. capital intensity (log)			-0.0053 (0.0042)	-0.0160 (0.0170)
F test of excluded IV		42.29		66.36
N of establishments	10854	10854	10854	10854
N. obs	732991	732991	732991	732991

Notes: Unit of analysis: establishment-occupation (2-digit PCS) pair. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Additional controls in all specification: occupation-establishment fixed effects, occupation-year dummies, sector-year dummies, region-year dummies, ETS-year dummies, initial decile of electricity share x year dummies.

# Wages, 4digit, main results

Table: Results for wages (4-digit occupations)

Dep var: hourly wage (log)	(1)	(2)	(3)
Energy price (log)	-0.0260 (0.0184)		
Energy price (log) x Carbon cont. of occupation (log)	-0.0221*** (0.00722)		
Dirty-energy price (log)		-0.0127 (0.0161)	-0.00311 (0.0239)
Dirty-energy price (log) x Carbon cont. of occupations (log)		-0.0198** (0.00819)	-0.0183** (0.00777)
Dirty-energy price (log) x Sectoral emission intensity (log)			-0.00630 (0.00959)
F test of excluded IV	200.1	136.1	39.64
N of establishments	13036	10338	10338
N. obs	2335130	2030497	2030497

Notes: Unit of analysis: establishment-occupation (4-digit PCS) pair. IV-FE model. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Additional controls in all specification: occupation-establishment fixed effects, occupation-year dummies, sector-year dummies, region-year dummies, ETS-year dummies, initial decile of electricity share x year dummies (except column 1).

# Wage, 4digit, by occupation

**Table:** Results for wages (4-digit occupations) - by occupational skill level

Dep var: average hourly wage (log)	Low-skill occ	Medium- skill occ	High-skill occ
Dirty-energy price (log)	-0.00437 (0.0385)	-0.0160 (0.0170)	-0.00235 (0.0282)
Dirty-energy price (log) x Carbon cont. of occupations (log)	-0.0195 (0.0287)	-0.0265*** (0.00878)	0.00801 (0.0204)
F test of excluded IV	139.4	139.7	97.96
N of establishments	9402	10321	9858
N. obs	196002	1286159	548336

Notes: Unit of analysis: establishment-occupation (4-digit PCS) pair. IV-FE model. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Additional controls in all specification: occupation-establishment fixed effects, occupation-year dummies, sector-year dummies, region-year dummies, ETS-year dummies, initial decile of electricity share x year dummies.

# Discrete policy change: the EU-ETS

- We also consider the introduction of the **EU-ETS** as a **discrete policy change** affecting the **price of fossil fuels** for treated establishments
- We follow the **standard approach** popularized by Calel and Dechezlepretre (2016) of **matching** treated establishment with non-treated ones with **similar characteristics**
  - ◊ **Matching variables** (measured in 2004, EACEI sample of that year):
    - ▶ Energy-related CO2 emissions (log)
    - ▶ Establishment size (dummy for 250+ employees FTE)
    - ▶ Shares of employment in HS and MS occupations
    - ▶ Average carbon content of occupations of employees in the establishment (log)
    - ▶ Sector dummies (2-digit) ⇒ Exact match on sector
  - ◊ **One nearest neighbour** with replacement and caliper (1/4 SD of propensity score), common support ⇒ 279 treated, 158 matched controls
- We estimate a **diff-in-diff** on the **matched sample**, interacting the treatment with the carbon content of occupations



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# Balancing (I)

**Table:** Balancing between treated and untreated establishments: matching variables

Matching variables (year 2004)	Treated	Untreated	t-test	Treated match	Untreated match	t-test
CO2 emissions (log)	4.95	0.92	38.17***	4.70	4.75	-0.36
Medium-big estab (>250 emp)	0.52	0.24	11.69***	0.48	0.52	-0.85
Share of high skill	0.112	0.123	-1.66*	0.113	0.118	-0.50
Share of middle skill	0.77	0.677	8.60***	0.760	0.770	-0.79
Average carbon cont. of occ. (log)	10.12	10.04	11.01***	10.12	10.12	-0.60

## Balancing (II)

**Table:** Balancing between treated and untreated establishments: other variables

Matching variables (year 2004)	Treated	Untreated	t-test	Treated match	Untreated match	t-test
Energy use (log)	7.14	3.19	39.52***	6.93	6.76	1.34
Employment FTE (log)	5.46	4.58	15.11***	5.34	5.46	-1.32
Electricity share	0.287	0.584	-19.93***	0.292	0.249	2.61**
Average wage (log)	10.16	9.98	13.93***	10.15	10.12	2.05**
Capital intensity of occ. (log)	4.52	4.49	8.84***	4.52	4.52	-0.15

# EU-ETS: employment, 2digit

Table: Effect of the EU-ETS - employment

Dep var: FTE empl (log)	(1)	(2)	(3)	(4)
ETS x Post (2005-2019)	-0.0152 (0.0405)		0.0216 (0.0401)	
Carbon content of occ. (log)	0.0434 (0.0330)	0.0559 (0.0341)		
ETS x Post (2005-2019) x Carbon content of occ. (log)	-0.0777*** (0.0200)		-0.0245 (0.0263)	
ETS x Phase 1 (2005-2007)		0.0110 (0.0300)		-0.00247 (0.0299)
ETS x Phase 2 (2008-2012)		-0.0376 (0.0541)		-0.00148 (0.0576)
ETS x Phase 3 (2013-2019)		0.00614 (0.0663)		0.0959 (0.0766)
ETS x Phase 1 (2005-2007) x Carbon content of occ. (log)		-0.00399 (0.0163)		-0.0338 (0.0277)
ETS x Phase 2 (2008-2012) x Carbon content of occ. (log)		-0.0854*** (0.0206)		-0.0379 (0.0312)
ETS x Phase 3 (2013-2019) x Carbon content of occ. (log)		-0.0838*** (0.0246)		0.0320 (0.0368)
Occ. (2-digit)-year FE			✓	✓
N. obs	99784	99784	99784	99784

Notes: Unit of analysis: establishment-occupation (2-digit) pair. Standard errors clustered by establishment in parenthesis. \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Fixed effect model on matched sample. Additional controls: sector-by-year dummies, region-by-year dummies. Matching based on propensity score.

## EU-ETS: wage, 2digit

Table: Effect of the EU-ETS - wage

Dep var: average hourly wage (log)	(1)	(2)	(3)	(4)
ETS x Post (2005-2019)	-0.0155* (0.00907)		-0.0177**	
Carbon content of occ. (log)	0.0798*** (0.0140)	0.0810*** (0.0147)		
ETS x Post (2005-2019) x Carbon content of occ. (log)	0.0173** (0.00804)		0.0160 (0.0118)	
ETS x Phase 1 (2005-2007)		-0.0138 (0.00891)		-0.0148* (0.00834)
ETS x Phase 2 (2008-2012)		-0.0130 (0.0107)		-0.0184* (0.0101)
ETS x Phase 3 (2013-2019)		-0.0215* (0.0125)		-0.0241* (0.0126)
ETS x Phase 1 (2005-2007) x Carbon content of occ. (log)		0.00219 (0.00790)		0.00467 (0.0104)
ETS x Phase 2 (2008-2012) x Carbon content of occ. (log)		0.00322*** (0.00921)		0.0246* (0.0134)
ETS x Phase 3 (2013-2019) x Carbon content of occ. (log)		0.00621 (0.00912)		0.00361 (0.0133)
Occ. (2-digit)-year FE			✓	✓
N. obs	68578	68578	68578	68578

Notes: Unit of analysis: establishment-occupation (2-digit) pair. Standard errors clustered by establishment in parenthesis. \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Fixed effect model on matched sample. Additional controls: sector-by-year dummies, region-by-year dummies. Matching based on propensity score.

# Preliminary quantification of energy prices on employment

*How to interpret our results? How many workers are at risk of displacement for carbon pricing?*

- We do not use weights as the EACEI sample is highly selected (see Marin and Vona, 2021).
- The estimated effect is a cross-elasticity of labour demand, but it is a “LATE” and does not account for compositional effects.
- Marin and Vona (2021) and Dussaux (2019) show that compositional effects can both mitigate and amplify employment effects:
  - ◊ Exit of carbon-intensive firms amplifies it;
  - ◊ Labour reallocation towards less-carbon intensive, more productive establishments (both within the firm and the industry) mitigates it;
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## Preliminary quantification (cnt.)

- Given these premises, our estimates may **slightly understate the risk of displacement**.
- We compute the share of workers at risk of displacement in high-carbon occupations as follows:

$$RiskDisplCC_t = \sum_{o=1}^{411} \frac{L_{ot}}{L_t} \times \frac{L_{o \in \text{manuf}, t}}{L_{ot}} \times (\mathbb{1}_{o \in \Theta(\hat{\beta}_o)} - \mathbb{1}_{o \in \text{green}}),$$

- $\frac{L_{ot}}{L_t} \times \frac{L_{o \in \text{manuf}, t}}{L_{ot}}$  captures the size of the occupations  $X$  the exposure to the treatment, i.e. working in manufacturing;
  - $(\mathbb{1}_{o \in \Theta} - \mathbb{1}_{o \in \text{green}})$  is the subset of occupations for which estimated energy price effects ( $\hat{\beta}_o$ ) are negative and stat. significant and not green.
- Approximating the set  $\Theta(\hat{\beta}_o)$  to contain the **last quartile of high-carbon occupations**, we obtain:  $RiskDisplCC_{2003-06} = 8.4\%$  and  $RiskDisplCC_{2015-18} = 5.7\%$

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# Mechanisms behind the wage effects

- **Induced compositional effects**, which are induced by carbon price shocks, vs. **loss of rents**, which are usually higher in carbon-intensive sectors:
  - ◇ **AKM model**: estimate estab.-, worker- and estab.-by-worker fixed effects for the **universe of French companies** and **specific sub-periods** (Babet et al., 2022).
  - ◇ **Re-estimate** the effect of the **carbon pricing shocks** on those FE aggregated at the occupation-by-establishment level.
- **Stayers vs. movers**: we know that, especially in rigid labour markets such as the French one, wage effects more likely to emerge for movers.
  - ◇ Replicate the analysis for **various types of movers**: across occupations, across sectors and across occupation-sectors.

## THANKS FOR YOUR ATTENTION

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Labour in the Low-Carbon Transition, new program at the FEEM: <https://www.feem.it/en/ricerca/programmi/labour-in-the-low-carbon-transition/>



# Top Occupations in terms of carbon content

**Table:** Top occupations in terms of CO2 emission intensity [back](#)

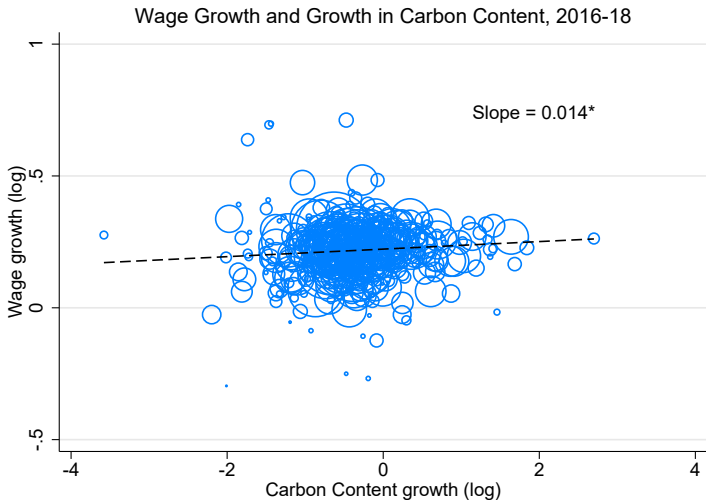
Job title	CO2/L init. (log)	Delta CO2/L (log)	Manuf Emp Sh
Capitaines timoniers navigation fluviale	13.20	-0.95	0.01
Officiers, cadres techniques marine marchande	13.13	-1.47	0.00
Matelots marine marchande	13.07	-1.28	0.01
Maitres d'équipage marine marchande, pêche	12.69	-1.18	0.02
Officiers, cadres navigants techniques aviation civile	12.57	-0.03	0.01
Hôtesse de l'air, stewards	12.51	0.11	0.00
Techniciens production, distribution ind. (énergie/eau/chauffage)	12.01	-0.16	0.07
Agents non-qualifiés services exploitation transports	11.91	-2.20	0.02
Mineurs qualifiés, autres ouvriers qualifiés extraction	11.80	-0.39	0.10
Pilotes d'installation lourde industries transformation	11.78	0.21	0.96
Bobiniers qualifiés	11.69	-3.58	0.92
Responsables commerciaux, administratifs transports voyageurs	11.63	0.12	0.00
Techniciens production, contrôle-qualité industries transformation	11.56	-0.73	0.90
Agents de maîtrise en fabrication	11.56	-0.46	0.93
Ouvriers qualifiés, autres ind. (énergie/eau/chauffage)	11.55	-0.19	0.14
Ingénieurs, cadres production distribution énergie/eau	11.55	0.07	0.07
Ouvriers production non-qualifiés imprimerie/presse/édition	11.54	-0.82	0.63
Conducteurs d'engin lourd de manœuvre	11.53	-0.41	0.13
Agents services commerciaux transports voyageurs	11.51	-0.74	0.00
Conducteurs d'engin lourd de levage	11.48	-0.41	0.32
<i>Unweighted averages occupation (top 20)</i>	12.01	-0.72	0.26
<i>Unweighted averages occupation (all)</i>	9.29	-0.35	0.23

# Top Occupations in terms of carbon content

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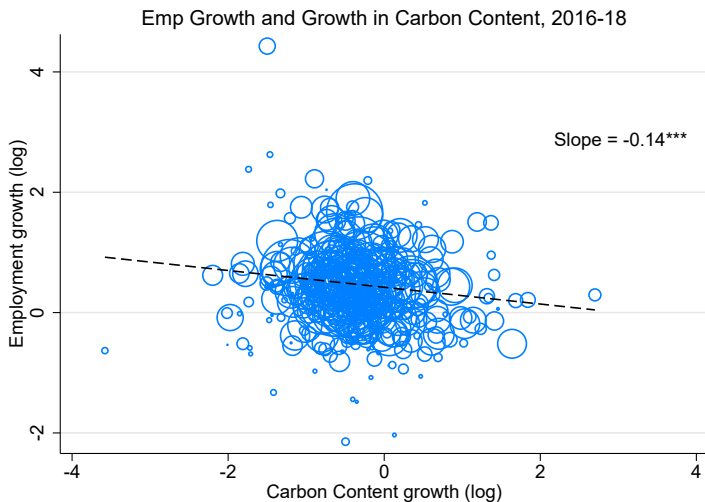
Job title	Imp China (log)	K/L (log)	Spat. Gini	Wage (log)
Capitaines timoniers navigation fluviale	-8.7	4.4	0.44	10.4
Officiers, cadres techniques marine marchande	-9.1	5.5	0.47	10.9
Matelots marine marchande	-7.4	5.4	0.47	10.2
Maitres d'équipage marine marchande, pêche	-7.7	5.1	0.47	10.3
Officiers, cadres navigants techniques aviation civile	-9.6	4.6	0.46	11.6
Hôtesse de l'air, stewards	-11.6	4.5	0.48	10.3
Techniciens production, distribution ind. (énergie/eau/chauffage)	-7.0	6.1	0.15	10.2
Agents non-qualifiés services exploitation transports	-7.7	5.3	0.21	9.9
Mineurs qualifiés, autres ouvriers qualifiés extraction	-6.0	5.0	0.17	10.0
Pilotes d'installation lourde industries transformation	-4.2	4.8	0.30	10.2
Bobiniers qualifiés	-3.2	3.5	0.34	10.0
Responsables commerciaux, administratifs transports voyageurs	-9.4	5.4	0.23	10.3
Techniciens production, contrôle-qualité industries transformation	-4.2	4.9	0.17	10.2
Agents de maîtrise en fabrication	-3.8	4.5	0.18	10.3
Ouvriers qualifiés, autres ind. (énergie/eau/chauffage)	-6.1	7.4	0.13	10.1
Ingénieurs, cadres production distribution énergie/eau	-6.8	6.8	0.19	10.9
Ouvriers production non-qualifiés imprimerie/presse/édition	-4.8	3.6	0.24	9.7
Conducteurs d'engin lourd de manœuvre	-6.3	4.6	0.20	10.1
Agents services commerciaux transports voyageurs	-10.1	4.8	0.17	10.0
Conducteurs d'engin lourd de levage	-5.3	4.2	0.18	10.1
<i>Unweighted averages occupation (top 20)</i>	-6.96	5.02	0.28	10.28
<i>Unweighted averages occupation (all)</i>	-7.01	4.05	0.20	10.18

# FACT IV: Growth-growth carbon content and wages



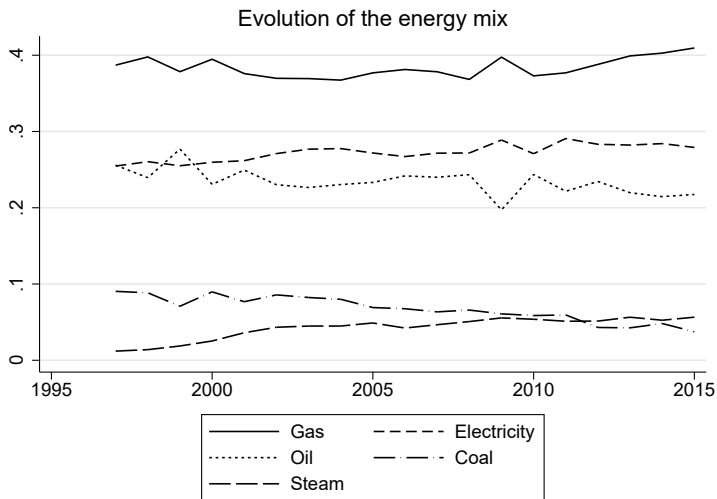
source: our elaborations on EACEI-DADS-JRC data [Back](#)

# FACT III: Growth-growth carbon content and emp.



source: our elaborations on EACEI-DADS-JRC data [Back](#)

# Energy mix



source: our elaborations on EACEI data [Back](#)

# Testing for parallel trends (from Marin and Vona, 2021)

**Table:** Tests for different trends for establishments with different initial energy mixes (unbalanced panel)

Dependent variable: Full-time equivalent employment (in log)				
F test: joint significance of electr share x time dummies	4.227	4.125	2.381	1.411
p-value	0.002	0.002	0.0493	<b>0.208</b>
F test: joint significance of gas share x time dummies	1.657	1.981	1.354	1.490
p-value	0.157	0.0946	<b>0.247</b>	<b>0.202</b>
Dependent variable: Average wage per employee FTE (euro, in log)				
F test: joint significance of electr share x time dummies	1.339	0.383	0.374	0.399
p-value	<b>0.253</b>	<b>0.821</b>	<b>0.827</b>	<b>0.810</b>
F test: joint significance of gas share x time dummies	1.583	1.236	1.090	1.151
p-value	<b>0.176</b>	<b>0.293</b>	<b>0.359</b>	<b>0.331</b>
Year dummies	Yes	-	-	-
Region x year dummies	-	Yes	Yes	Yes
Sector (2-digit) x year dummies	-	-	Yes	Yes
Additional controls	-	-	-	Yes

Notes: Fixed effect model. Standard errors clustered by establishment in parenthesis. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Sample: establishments of the 1997 sample in years 2001 (included establishments should be observed at least twice). Regressions include the interaction between initial shares (of gas and electricity, respectively) and year dummies. Gas includes natural gas, butane-propane, other gases.  $N=32676$  (28783 for CO2).